Problem Set 2

Econ 211C

$$f_{Y_i}(y_i|k,\lambda) = \frac{k}{\lambda} \left(\frac{y_i}{\lambda}\right)^{k-1} \exp\left\{-\left(\frac{y_i}{\lambda}\right)^k\right\}.$$

(a) (10 points) Derive the log likelihood, $\ell(k, \lambda | \boldsymbol{y})$.

Solution: The likelihood function is equivalent to the joint density:

$$\mathcal{L}(k,\lambda|\mathbf{y}) = f_{\mathbf{Y}}(\mathbf{y}|k,\lambda)$$

$$= \prod_{i=1}^{n} f_{Y_i}(y_i|k,\lambda)$$

$$= \prod_{i=1}^{n} \frac{k}{\lambda} \left(\frac{y_i}{\lambda}\right)^{k-1} \exp\left\{-\left(\frac{y_i}{\lambda}\right)^k\right\}$$

$$= k^n \lambda^{-nk} \exp\left\{-\lambda^{-k} \sum_{i=1}^{n} y_i^k\right\} \prod_{i=1}^{n} y_i^{k-1}.$$

Thus the log likelihood is

$$\ell(k, \lambda | \boldsymbol{y}) = n \log(k) - nk \log(\lambda) - \lambda^{-k} \sum_{i=1}^{n} y_i^k + (k-1) \sum_{i=1}^{n} \log(y_i).$$

(b) (10 points) Derive the maximum likelihood estimators, \hat{k} and $\hat{\lambda}$.

Solution: We begin by taking derivatives of the log likelihood function:

$$\frac{\partial \ell}{\partial k} = \frac{n}{k} - n \log(\lambda) + \log(\lambda) \lambda^{-k} \sum_{i=1}^{n} y_i^k - \lambda^{-k} \sum_{i=1}^{n} \log(y_i) y_i^k + \sum_{i=1}^{n} \log(y_i)$$
 (1)

$$\frac{\partial \ell}{\partial \lambda} = -\frac{nk}{\lambda} + k\lambda^{-(k+1)} \sum_{i=1}^{n} y_i^k. \tag{2}$$

Beginning with Equation (2), the MLEs, \hat{k} and $\hat{\lambda}$, are the values such that

$$\frac{\partial \ell}{\partial \lambda} \Big|_{\hat{k},\hat{\lambda}} = 0$$

$$\implies n\hat{k}\hat{\lambda}^{-1} = \hat{k}\hat{\lambda}^{-(\hat{k}+1)} \sum_{i=1}^{n} y_i^{\hat{k}}$$

$$\implies n\hat{\lambda}^{\hat{k}} = \sum_{i=1}^{n} y_i^{\hat{k}}$$

$$\implies \hat{\lambda} = \left(\frac{1}{n} \sum_{i=1}^{n} y_i^{\hat{k}}\right)^{\frac{1}{\hat{k}}}.$$
(3)

Substituting Equation (3) into Equation (1),

$$\left. \frac{\partial \ell}{\partial k} \right|_{\hat{k}, \hat{\lambda}} = \frac{n}{\hat{k}} - \frac{n \sum_{i=1}^{n} \log(y_i) y_i^{\hat{k}}}{\sum_{i=1}^{n} y_i^{\hat{k}}} + \sum_{i=1}^{n} \log(y_i) = 0.$$
 (4)

Equation (4) defines a unique value for \hat{k} , but it must be determined numerically since it can't be solved analytically.

(c) (15 points) Derive the information matrix. What is the observed information matrix? Given estimates, \hat{k} and $\hat{\lambda}$, what would approximations of the variances of the estimates be?

Solution: Beginning with Equations (1) and (2), we derive the second derivatives of the log likelihood function:

$$\begin{split} \frac{\partial^2 \ell}{\partial k^2} &= -\frac{n}{k^2} - \log(\lambda)^2 \lambda^{-k} \sum_{i=1}^n y_i^k + 2\log(\lambda) \lambda^{-k} \sum_{i=1}^n \log(y_i) y_i^k - \lambda^{-k} \sum_{i=1}^n \log(y_i)^2 y_i^k \\ \frac{\partial^2 \ell}{\partial k \partial \lambda} &= \frac{\partial^2 \ell}{\partial \lambda \partial k} = -\frac{n}{k} + \lambda^{-(k+1)} \left[(1 - k \log(\lambda)) \sum_{i=1}^n y_i^k + k \sum_{i=1}^n \log(y_i) y_i^k \right] \\ \frac{\partial^2 \ell}{\partial \lambda^2} &= nk \lambda^{-2} - k(k+1) \lambda^{-(k+2)} \sum_{i=1}^n y_i^k. \end{split}$$

The Hessian is the matrix

$$\mathcal{H}(k,\lambda) = \begin{bmatrix} \frac{\partial^2 \ell}{\partial k^2} & \frac{\partial^2 \ell}{\partial k \partial \lambda} \\ \frac{\partial^2 \ell}{\partial \lambda \partial k} & \frac{\partial^2 \ell}{\partial \lambda^2} \end{bmatrix},$$

and the information matrix is $\mathcal{I}(k,\lambda) = -\mathbb{E}\left[\mathcal{H}(k,\lambda)^{-1}\right]$. The observed information matrix is simply $-\mathcal{H}(k,\lambda)^{-1}$ and approximations of the variances of \hat{k} and $\hat{\lambda}$ are given by the diagonal elements of $-\mathcal{H}(\hat{k},\hat{\lambda})^{-1}$, where we evaluate the inverse hessian at the estimated values, \hat{k} and $\hat{\lambda}$.

Consider an AR(2) process

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-1} + \varepsilon_t$$

where $\varepsilon_t \stackrel{i.i.d}{\sim} \mathcal{N}(0,1)$ and where $\boldsymbol{\phi} = (\phi_1, \phi_2)' = (1.3, -0.41)'$.

(a) (25 points) Simulate 30 observations from this process and compute the least-squares estimates for three regressions:

$$Y_{t} = \phi_{1}Y_{t-1} + \varepsilon_{t}$$

$$Y_{t} = \phi_{1}Y_{t-1} + \phi_{2}Y_{t-2} + \varepsilon_{t}$$

$$Y_{t} = \phi_{1}Y_{t-1} + \phi_{2}Y_{t-2} + \phi_{3}Y_{t-3} + \varepsilon_{t}.$$

Repeat the simulation/estimation 1000 times and report the means and standard deviations of each set of estimates in a table. Include your R code with your solution.

Solution: The code to generate the results for parts (a), (b) and (c) is shown below:

Parameters

N = 1000

nSim = 100000

phi1 = 1.3

phi2 = -0.41

reg1 = NULL

reg2 = NULL

reg3 = NULL

for(ix in 1:nSim){

```
# Simulate
eps = rnorm(N+2)
Y = rep(0, N+2)
for(jx in 3:(N+2)){
Y[jx] = phi1*Y[jx-1]+phi2*Y[jx-2] + eps[jx]
}
Y = Y[3:(N+2)]
# Estimate
reg1 = rbind(reg1, lm(Y[2:N]^Y[1:(N-1)]-1)$coef)
reg2 = rbind(reg2, lm(Y[3:N]^{Y}[2:(N-1)]+Y[1:(N-2)]-1)$coef)
reg3 = rbind(reg3,lm(Y[4:N]^Y[3:(N-1)]+Y[2:(N-2)]+Y[1:(N-3)]-1)$coef)
}
# Compute moments
signif(apply(reg1,2,mean),4)
signif(apply(reg1,2,sd),4)
signif(apply(reg2,2,mean),4)
signif(apply(reg2,2,sd),4)
signif(apply(reg3,2,mean),4)
signif(apply(reg3,2,sd),4)
```

The only modification that must be made for each part is on the second line, modifying the value of N. The corresponding results are reported in the table below.

\overline{N}	ϕ_1	ϕ_2	ϕ_3
		AR(1)	
30	0.9089 (0.06614)	NA	NA
1000	0.9215 (0.008059)	NA	NA
		AR(2)	
30	1.265 (0.1826)	-0.4090 (0.1789)	NA
1000	1.299 (0.02910)	-0.4105 (0.02889)	NA
		AR(3)	
30	1.264 (0.2007)	-0.4110 (0.2918)	0.0004006 (0.1898)
1000	1.299 (0.03196)	-0.4105 (0.05125)	$0.00001385 \ (0.03263)$

- (b) (5 points) Repeat part (a), simulating N = 1000 observations instead of N = 30 observations at each iteration. Report the estimates in the same table as part (a).
- (c) (5 points) Repeat part (b) with N = 100,000.

Solution: The code to load data and fit ARMA models is shown below.

```
library(quantmod)
getSymbols('XIV',from='2014-04-25',to='2015-04-24')
rets = dailyReturn(XIV)
auto.arima(rets)
```

In this case, I used the auto.arima fuction from the forecast package in R to sequentially search over ARMA(p,q) models for $p=0,\ldots,5$ and $q=0,\ldots,5$. The function can use a variety of information criteria (penalized log likelihoods) to select the best fitting model, which balances the objectives of maximizing the likelihood and maintaining a low order of parameterization. Using BIC (Bayeisan Information Criterion), auto.arima selected an ARMA(0,0) with zero mean as the best fitting model. This suggests that daily XIV returns are white noise, or that daily XIV prices are a random walk.